

# Effect of a conditional cash transfer programme on childhood mortality: a nationwide analysis of Brazilian municipalities



Davide Rasella, Rosana Aquino, Carlos A T Santos, Rômulo Paes-Sousa, Mauricio L Barreto

## Summary

**Background** In the past 15 years, Brazil has undergone notable social and public health changes, including a large reduction in child mortality. The *Bolsa Família* Programme (BFP) is a widespread conditional cash transfer programme, launched in 2003, which transfers cash to poor households (maximum income US\$70 per person a month) when they comply with conditions related to health and education. Transfers range from \$18 to \$175 per month, depending on the income and composition of the family. We aimed to assess the effect of the BFP on deaths of children younger than 5 years (under-5), overall and resulting from specific causes associated with poverty: malnutrition, diarrhoea, and lower respiratory infections.

**Methods** The study had a mixed ecological design. It covered the period from 2004–09 and included 2853 (of 5565) municipalities with death and livebirth statistics of adequate quality. We used government sources to calculate all-cause under-5 mortality rates and under-5 mortality rates for selected causes. BFP coverage was classified as low (0·0–17·1%), intermediate (17·2–32·0%), high (>32·0%), or consolidated (>32·0% and target population coverage ≥100% for at least 4 years). We did multivariable regression analyses of panel data with fixed-effects negative binomial models, adjusted for relevant social and economic covariates, and for the effect of the largest primary health-care scheme in the country (Family Health Programme).

**Findings** Under-5 mortality rate, overall and resulting from poverty-related causes, decreased as BFP coverage increased. The rate ratios (RR) for the effect of the BFP on overall under-5 mortality rate were 0·94 (95% CI 0·92–0·96) for intermediate coverage, 0·88 (0·85–0·91) for high coverage, and 0·83 (0·79–0·88) for consolidated coverage. The effect of consolidated BFP coverage was highest on under-5 mortality resulting from malnutrition (RR 0·35; 95% CI 0·24–0·50) and diarrhoea (0·47; 0·37–0·61).

**Interpretation** A conditional cash transfer programme can greatly contribute to a decrease in childhood mortality overall, and in particular for deaths attributable to poverty-related causes such as malnutrition and diarrhoea, in a large middle-income country such as Brazil.

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## Introduction

Conditional cash transfer programmes are interventions that transfer cash from governments to poor households with the requirement that parents comply with specific conditions (or conditionalities), usually focused on health and education for their children.<sup>1</sup> The transfer of benefits aims to promptly alleviate poverty and the conditions encourage use of existing health and education services. The first conditional cash transfer programmes were implemented in the late 1990s in Mexico and Brazil, spreading rapidly to various countries worldwide, becoming an important strategy for alleviation of poverty and reduction of inequalities in low-income and middle-income countries.<sup>1,2</sup>

In Brazil, the *Bolsa Família* programme (Family Allowance, BFP), launched in 2003, merged four pre-existing national social programmes into one unique expanded programme.<sup>3</sup> The BFP is the world's largest conditional cash transfer programme, and its coverage has expanded greatly in the past 10 years. It reached all 5565 Brazilian municipalities and enrolled 13·4 million

families in 2011, with a total budget of US\$11·2 billion.<sup>4</sup> The cash transfers are intended for extremely poor families (with an income of less than \$35 per person per month) and for other families deemed poor (with an income of between \$35 and \$70 per person per month) when they include children up to 17 years of age or pregnant or lactating women.<sup>5</sup> Poor families receive about \$18 for each pregnant woman, child, or adolescent up to 17 years of age (with an upper limit for each category), whereas extremely poor families, besides receiving the same benefits, receive an additional contribution of \$35 irrespective of the composition of the family. According to these criteria, benefits can range from \$18 to a maximum of \$175 per month. The mother (when present) must receive the monthly payment on behalf of the whole family.

A family enrolled in the BFP has to comply with specific education and health-related conditions. To meet the health conditions children younger than 7 years must be vaccinated according to the Brazilian immunisation programme schedule and must comply with health check-ups and growth monitoring according to Ministry

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See [Comment](#) page 7

Instituto de Saúde Coletiva,  
Federal University of Bahia,  
Salvador, Bahia, Brazil

(D Rasella PhD, R Aquino MD,  
C A T Santos PhD,  
Prof M L Barreto MD);  
Department of Exact Sciences,  
State University of Feira de  
Santana, Feira de Santana,  
Bahia, Brazil (C A T Santos);  
Institute of Development  
Studies, University of Sussex,  
Brighton, UK

(R Paes-Sousa MD); and *Ciência,  
Tecnologia e Inovação em  
Saúde, INCT-CITECS*, Salvador,  
Bahia, Brazil (Prof M L Barreto)

Correspondence to:

Prof Mauricio Barreto, Instituto  
de Saúde Coletiva, Universidade  
Federal de Bahia,  
Canela, 41110-040,  
Salvador, Bahia, Brazil  
[mauricio@ufba.br](mailto:mauricio@ufba.br)

of Health guidelines, with a frequency from one to seven times per year, depending on a child's age. Pregnant and lactating women must attend scheduled prenatal and postnatal visits and health and nutritional educational activities. When possible, health-related conditions should be met using the facilities of the main primary health care programme in Brazil, the *Programa Saúde da Família* (Family Health Programme, FHP).<sup>6</sup> The FHP is another large-scale national programme, implemented over the past several years. By 2011, it reached 94% of municipalities, covering 53% of the Brazilian population.<sup>7</sup> FHP aims to broaden access to public health services, especially in deprived areas, by offering free, community-based health care.<sup>8</sup>

Brazil is characterised by large social inequalities, and it has undergone substantial health and social changes in the past 15 years, including a large reduction in deaths of children younger than 5 years (under-5), enabling the country to reach the fourth Millennium Development Goal.<sup>9,10</sup> FHP is one of the components that has brought about the substantial decrease in under-5 mortality.<sup>11,12</sup> We postulate that the BFP should reduce childhood mortality by acting on social determinants of health and by stimulating health care through its conditions. Previous studies have reported the effectiveness of BFP in reducing child malnutrition,<sup>13,14</sup> but no studies have addressed its effect on childhood morbidity or mortality. Therefore, the objective of the present study was to assess the effect of the BFP on under-5 mortality rates in Brazilian municipalities, focussing on causes of mortality associated with poverty (such as malnutrition, diarrhoea, and lower respiratory infections) and on some of the potential intermediate mechanisms (such as vaccination, prenatal care, and admission to hospital).

## Methods

### Study design

This study has a mixed ecological design, combining an ecological multiple-group design with a time-trend design. The municipality is the unit of analysis. We created a longitudinal dataset from several databases for the years 2004–09. From the 5565 Brazilian municipalities, we selected a subset that had adequate vital statistics (death and livebirth registration) during the first years of the period under study (2004–06; we assumed constant adequacy for the remaining years because of improvements in collection of vital information).<sup>15</sup> We assessed adequacy of mortality information according to a validated multi-dimensional criterion,<sup>16</sup> which took into account the value of the age-standardised mortality rate of the municipality, the ratio between registered and estimated birth rates, the percentage of poorly defined deaths, and the mean deviation of the previous two parameters for 2004–06. We obtained mortality rates by direct calculation—the number of deaths of children younger than 5 years per 1000 livebirths. Groups of selected causes of mortality and admission to hospital were created by aggregation of

categories from the International Classification of Diseases, 10th revision:<sup>17</sup> diarrhoeal diseases (A00, A01, A03, A04, A06–09), malnutrition (E40–46), lower respiratory infections (J10–18, J20–22), and external causes (V01–98). Mortality attributed to external causes (which includes transport accidents, homicides, and accidental injuries) was included as a control because no effect from either of the programmes was expected. Rates of under-5 admission to hospital were also obtained by direct calculation. A vaccination coverage index for children younger than 1 year was created with areas dichotomised into those where coverage of three main vaccines (measles, oral polio, and diphtheria-pertussis-tetanus [DPT]) was higher than 95% and those with lower coverage.

For the BFP, it is possible to conceive of two indicators of coverage. The first is coverage of the target population, calculated as the number of families enrolled in the BFP in a municipality divided by the number of eligible families (according to BFP criteria) in the same municipality.<sup>4</sup> The second is coverage of the total population, calculated as the number of individuals enrolled in the BFP (obtained by multiplying the number of beneficiary families by the average family size) divided by the total population of the same municipality. All models have been fitted using these two indicators (appendix p 4). We have also created a coverage indicator combining both indicators. The categories for this BFP coverage indicator were: low (BFP coverage of the total population of the municipality from 0·0% to 17·1%), intermediate (17·2–32·0%), high (>32·0%), and consolidated (BFP coverage of the total population of the municipality >32·0% and, at the same time, BFP coverage of the target population  $\geq$ 100% for at least the previous 4 years). Because of the absence in the scientific literature of previous reference values, the cutoffs used for the categorisation (17·1% and 32·0%) represented the tertiles of the distribution of BFP coverage of the total population. This indicator, adjusted in the models for the percentage of the target population in the municipality, enabled us to capture the effect of programme duration and the effect of possible programme externalities (ie, positive spillover effects on programme-ineligible inhabitants) in the municipality.<sup>2</sup> We calculated yearly coverage of the FHP as the ratio of the total number of individuals registered in this programme to the population of the municipality, and it was categorised, for comparability reasons, as in previous studies:<sup>11,12,15</sup> without FHP, incipient (<30·0% of the population), intermediate (30·0–69·9% or  $\geq$ 70·0% for less than the previous 4 years), and consolidated ( $\geq$ 70·0% for at least the previous 4 years). We selected a set of covariates recognised as determinants of under-5 mortality (all-cause and cause-specific) that were available at the municipality level: monthly income per person, proportion of total municipality population eligible for BFP, prevalence of illiteracy among individuals older than 15 years, percentage of individuals living in households with inadequate sanitation (inadequate water supply, sewers, and garbage collection), total fertility rate, and

See Online for appendix

overall rate of admissions to hospital in the municipality. We dichotomised the covariables according to the median value or, when available (as in the case of illiteracy and fertility rates), reference values.

### Data sources

The data used in this study were collected from different information systems. The data sources provided by the Ministry of Health were: mortality information system (under-5 deaths), primary care information system (FHP coverage), information system on livebirths (livebirths), and outpatient information system (admissions to hospital).<sup>7</sup> We used the Ministry of Social Development databases to calculate BFP coverage,<sup>8</sup> and we used data from the Brazilian Institute of Geography and Statistics for socioeconomic variables.<sup>18</sup> Because data for the covariates were obtained from the 2000 and 2010 national census databases, we calculated the annual values from 2004–09 by linear interpolation.

### Statistical analyses

We used conditional negative binomial regression models for panel data—consisting of a relevant number of units of analysis with repeated observations over time—with fixed-effects specification. As explained in detail in the appendix pp 7–9, to verify whether these models were actually removing the individual fixed effects,<sup>19</sup> we fitted models with different specifications from our dataset, including unconditional negative binomial regression models and conditional Poisson regressions with robust SEs. Conditional fixed-effects negative binomial regression models were shown to be the most appropriate for our analysis. The fixed-effects models, as with any other

longitudinal or panel data models, include a second term to control for characteristics of the unit of analysis that are constant during the study period and that have not been included in the model as confounding variables, such as some geographical, historical, or sociocultural aspects of each municipality. We chose the fixed-effects model specification on the basis of the Hausman test and because it is the most appropriate test for assessment of effects in interventions with panel data.<sup>20,21</sup> We did goodness of fit tests with Akaike information criterion and Bayesian information criterion estimates.<sup>19</sup> We fitted the same models with continuous or categorised variables (appendix pp 3–4). Whereas continuous variables allow estimation of the average strength of an association along the entire range of values for a variable, categorised variables give a measure of effect that is easier to interpret, comparing defined ranges of values. Moreover, use of different levels of coverage allows verification of the existence of a gradient of effect, related—in our study—to different degrees of implementation of the interventions.<sup>11,12,15</sup> To assess the association between BFP or FHP coverage and mortality rates, we calculated mortality rate ratios (RRs), both crude and adjusted for covariates, using municipalities with the lowest coverage as the reference category.

To detect any interaction between the BFP and FHP with regard to the reduction of all-cause and cause-specific under-5 mortality, we created a product term between the BFP and FHP coverage—both dichotomised as consolidated or not consolidated—and fitted models with the same specification as the previous ones but with this term representing the interaction between the two programmes. We did a sensitivity analysis with data from all Brazilian municipalities irrespective of quality of vital information.

|   | 2004         | 2005         | 2006         | 2007         | 2008         | 2009         | Percentage change 2004–09 |
|---|--------------|--------------|--------------|--------------|--------------|--------------|---------------------------|
| Mortality rate for children younger than 5 years (per 1000 livebirths)    |              |              |              |              |              |              |                           |
| Overall   | 21.7 (14.7)  | 20.3 (14.5)  | 20.1 (14.6)  | 19.4 (14.8)  | 18.6 (15.9)  | 17.5 (14.7)  | -19.4%                    |
| For diarrhoeal diseases   | 0.95 (2.93)  | 0.86 (2.54)  | 0.83 (2.67)  | 0.55 (2.02)  | 0.49 (1.96)  | 0.51 (2.46)  | -46.3%                    |
| For malnutrition  | 0.55 (2.33)  | 0.48 (2.24)  | 0.36 (1.70)  | 0.30 (2.53)  | 0.20 (1.26)  | 0.23 (1.54)  | -58.2%                    |
| For lower respiratory infections  | 1.15 (3.30)  | 0.96 (2.72)  | 1.07 (2.84)  | 0.95 (2.91)  | 0.98 (3.85)  | 0.84 (2.84)  | -27.0%                    |
| For external causes   | 1.23 (3.29)  | 1.16 (3.14)  | 1.06 (3.17)  | 1.16 (3.80)  | 1.07 (3.70)  | 1.01 (3.71)  | -17.9%                    |
| BFP coverage of the municipality population (%)                           | 17.3% (12.1) | 23.0% (14.0) | 28.1% (17.2) | 27.8% (17.8) | 25.2% (16.7) | 28.3% (17.5) | 63.6%                     |
| FHP coverage of the municipality population (%)                           | 62.7% (36.7) | 67.8% (34.8) | 71.0% (33.4) | 73.9% (32.4) | 74.4% (31.3) | 75.0% (30.9) | 19.6%                     |
| Income per person (monthly, in BR\$)                                      | 310 (126)    | 339 (135)    | 368 (145)    | 396 (154)    | 425 (164)    | 454 (147)    | 46.5%                     |
| Proportion of BFP eligible population in the municipality                 | 27.9% (16.5) | 27.8% (16.7) | 27.8% (16.8) | 27.7% (16.9) | 26.5% (15.5) | 26.3% (15.5) | -5.7%                     |
| Proportion of individuals living in households with inadequate sanitation | 22.9% (16.4) | 21.7% (15.8) | 20.5% (15.2) | 19.3% (14.7) | 18.2% (14.3) | 17.0% (13.9) | -25.8%                    |
| Proportion of individuals older than 15 years who are illiterate          | 16.9% (10.3) | 16.4% (10.0) | 15.9% (9.8)  | 15.4% (9.6)  | 14.9% (9.3)  | 14.4% (9.1)  | -14.8%                    |
| Total fertility rate  | 2.31 (0.62)  | 2.27 (0.63)  | 2.20 (0.64)  | 2.14 (0.65)  | 2.07 (0.65)  | 2.01 (0.67)  | -13.0%                    |
| Rate of admissions to hospital (per 100 inhabitants)                      | 4.88 (4.47)  | 4.69 (4.34)  | 4.58 (4.39)  | 4.46 (4.11)  | 4.02 (4.11)  | 4.04 (4.23)  | -17.2%                    |

Data are mean (SD). Causes of death were defined according to the International Classification of Diseases, 10th revision:<sup>27</sup> diarrhoeal diseases (A00, A01, A03, A04, A06–09), malnutrition (E40–46), lower respiratory infections (J10–18, J20–22), and external causes (V01–98). Rate of admission to hospital was calculated as the number of admissions to hospital for all ages and all causes of one municipality divided by the total population of the same municipality and multiplied by 100. BFP=Bolsa Família Programme. FHP=Family Health Programme.

**Table 1: Mortality rates and variables for selected municipalities (N=2853)**

|   | BFP models       |                  | FHP models       |                  | FHP and BFP (adjusted) |
|---|------------------|------------------|------------------|------------------|------------------------|
|   | Crude            | Adjusted         | Crude            | Adjusted         |                        |
| <b>BFP population coverage</b>  |                  |                  |                  |                  |                        |
| Low (0.0–17.1%)   | 1.00             | 1.00             | ..               | ..               | 1.00                   |
| Intermediate (17.2–32.0%)   | 0.91 (0.90–0.93) | 0.93 (0.91–0.95) | ..               | ..               | 0.94 (0.92–0.96)       |
| High (>32.0%)   | 0.82 (0.80–0.85) | 0.86 (0.83–0.89) | ..               | ..               | 0.88 (0.85–0.91)       |
| Consolidated (>32.0% and TPC ≥100% for at least 4 years)                          | 0.76 (0.72–0.80) | 0.81 (0.76–0.85) | ..               | ..               | 0.83 (0.79–0.88)       |
| <b>FHP municipality population coverage</b>                                       |                  |                  |                  |                  |                        |
| No FHP (0.0%)   | ..               | ..               | 1.00             | 1.00             | 1.00                   |
| Incipient (<30%)  | ..               | ..               | 0.97 (0.92–1.02) | 0.98 (0.94–1.03) | 0.99 (0.94–1.04)       |
| Intermediate (≥30%)   | ..               | ..               | 0.89 (0.85–0.93) | 0.91 (0.87–0.96) | 0.93 (0.88–0.97)       |
| Consolidated (≥70% and implemented for at least 4 years)                          | ..               | ..               | 0.81 (0.77–0.86) | 0.85 (0.80–0.90) | 0.88 (0.83–0.93)       |
| Income per person (monthly, >BR\$380)*  | ..               | 0.94 (0.92–0.97) | ..               | 0.93 (0.91–0.96) | 0.95 (0.92–0.97)       |
| Proportion of municipality population eligible for BFP* >22.4%                    | ..               | 1.07 (1.02–1.11) | ..               | 1.10 (1.06–1.15) | 1.07 (1.03–1.12)       |
| Proportion of individuals living in households with inadequate sanitation* <16.7% | ..               | 1.10 (1.05–1.15) | ..               | 1.11 (1.06–1.16) | 1.10 (1.05–1.15)       |
| Proportion of individuals older than 15 years who are illiterate† >11.1%          | ..               | 1.04 (1.00–1.09) | ..               | 1.05 (1.01–1.10) | 1.04 (1.00–1.08)       |
| Total fertility rate‡ >2.32   | ..               | 1.08 (1.04–1.11) | ..               | 1.08 (1.05–1.12) | 1.07 (1.03–1.10)       |
| Rate of admission to hospital (per 100 inhabitants)* >4.27                        | ..               | 1.02 (0.99–1.04) | ..               | 1.02 (0.99–1.04) | 1.01 (0.99–1.04)       |
| Number of observations  | 17 118           | 17 118           | 17 118           | 17 118           | 17 118                 |
| Number of municipalities  | 2853             | 2853             | 2853             | 2853             | 2853                   |

Data are rate ratio (95% CI) unless otherwise specified. TPC=target population coverage. \*Cutoff is median value. †Cutoff taken from Rasella and colleagues, 2010.<sup>12</sup>

**Table 2: Fixed-effect negative binomial models for association between under-5 mortality rates and Bolsa Familia Programme (BFP) and Family Health Programme (FHP) coverage**

We used Stata (version 12.0) for database processing and analysis.

**Role of the funding source**

The sponsor of the study had no role in study design, data collection, data analysis, data interpretation, or writing of the report. The corresponding author had full access to all the data in the study and had final responsibility for the decision to submit for publication.

**Results**

The criteria for adequate death and livebirth registration were met by 2906 municipalities. Of these, 2853 (51% of all Brazilian municipalities) had data available for all covariates and were included in our analysis. From 2004–09, the mean under-5 mortality rate decreased by 19.4% in the studied municipalities, and among the selected causes, the greatest decrease was associated with malnutrition (58.2%; table 1). Under-5 mortality associated with external causes decreased by 17.9%. Mean BFP coverage in the municipalities exhibited some yearly variation during the study period, reaching a peak in 2009 with 28.3% coverage. Mean FHP coverage in the municipalities continually increased, reaching 75.0% in 2009. Socioeconomic conditions improved during the study period, with the mean monthly income per person increasing by 46.5% and the percentage of individuals living in households with inadequate sanitation decreasing by 25.8% (table 1).

Table 2 shows the crude and adjusted associations of under-5 mortality rate with BFP and FHP municipal coverage levels. In the analysis, both measures of BFP and FHP coverage exhibited a significant dose–response association with decreasing under-5 mortality rate, even after the adjustment for socioeconomic and demographic covariates. We identified much the same results in models for which all the variables were imputed as continuous (appendix p 3). Table 3 shows the effect of BFP and FHP municipal coverage on selected causes of under-5 mortality. Both interventions had an effect on all selected causes except for external causes, which were used as a control. The strongest effect of the BFP was on under-5 mortality resulting from malnutrition, whereas the FHP was associated with the greatest reduction in diarrhoeal diseases and lower respiratory infections (table 3).

As shown in table 4, in multivariable models that controlled for FHP coverage and relevant covariates, the increase in BFP coverage increased vaccination coverage for measles, polio, and DPT, reduced the number of pregnant women who delivered without receiving any prenatal care, and reduced rates of under-5 admissions to hospital in a manner much the same as for the reduction in mortality rates, having the strongest effect on malnutrition and no effect on external causes.

When we included an interaction term between BFP and FHP in the models, this term was negatively associated with all mortality rates, both overall or for specific causes, but the association was significant only

|   | Diarrhoeal diseases | Malnutrition     | Lower respiratory infections | External causes  |
|---|---------------------|------------------|------------------------------|------------------|
| BFP municipality population coverage                    |                     |                  |                              |                  |
| Low (0.0–17.1%)   | 1.00                | 1.00             | 1.00                         | 1.00             |
| Intermediate (17.2–32.0%)                               | 0.83 (0.74–0.92)    | 0.66 (0.57–0.77) | 0.96 (0.88–1.05)             | 1.03 (0.95–1.13) |
| High (>32.0%)   | 0.68 (0.59–0.80)    | 0.54 (0.44–0.67) | 0.94 (0.82–1.07)             | 0.92 (0.79–1.06) |
| Consolidated (>32.0% and TPC ≥100 for at least 4 years) | 0.47 (0.37–0.61)    | 0.35 (0.24–0.50) | 0.80 (0.64–0.99)             | 0.92 (0.72–1.16) |
| FHP municipality population coverage                    |                     |                  |                              |                  |
| No FHP (0.0%)   | 1.00                | 1.00             | 1.00                         | 1.00             |
| Incipient (<30%)  | 0.90 (0.67–1.17)    | 0.88 (0.60–1.29) | 0.83 (0.68–1.00)             | 0.95 (0.79–1.14) |
| Intermediate (≥30%)                                     | 0.71 (0.54–0.93)    | 0.72 (0.49–1.07) | 0.71 (0.58–0.86)             | 0.87 (0.72–1.05) |
| Consolidate (≥70% and implemented for at least 4 years) | 0.53 (0.39–0.71)    | 0.59 (0.38–0.91) | 0.70 (0.56–0.87)             | 0.87 (0.70–1.08) |
| Number of observations                                  | 7356                | 5124             | 9894                         | 10776            |
| Number of municipalities                                | 1226                | 854              | 1649                         | 1796             |

Data are rate ratio (95% CI) unless otherwise specified. Models adjusted for income per person, proportion of municipality population eligible for BFP, proportion of individuals living in households with inadequate sanitation, proportion of individuals older than 15 years who are illiterate, total fertility rate, and rate of admissions to hospital.

**Table 3: Fixed-effect negative binomial models for adjusted associations between *Bolsa Familia* Programme (BFP) and Family Health Programme (FHP) coverage and under-5 mortality rates for some relevant groups of causes**

|   | Measles, polio, and DPT vaccine coverage over 95% among children younger than 1 year OR* (95% CI) | Proportion of pregnant women with no prenatal visits at the moment of delivery RR† (95% CI) | Under-5 rate of admission to hospital RR‡ (95% CI) | Under-5 rate of admission to hospital for diarrhoeal diseases RR‡ (95% CI) | Under-5 rate of admission to hospital for malnutrition RR‡ (95% CI) | Under-5 rate of admission to hospital for lower respiratory infections RR‡ (95% CI) | Under-5 rate of admission to hospital for external causes RR‡ (95% CI) |
|---|---|---|--|--|---|---|--|
| BFP municipality population coverage                    |   |   |  |  |   |   |  |
| Low (0.0–17.1%)   | 1.00  | 1.00  | 1.00   | 1.00   | 1.00  | 1.00  | 1.00   |
| Intermediate (17.2–32.0%)                               | 1.47 (1.29–1.68)  | 0.85 (0.83–0.87)  | 0.96 (0.95–0.97)                                   | 0.86 (0.84–0.87)   | 0.82 (0.76–0.87)  | 0.95 (0.93–0.97)  | 1.30 (0.85–1.99)   |
| High (>32.0%)   | 1.98 (1.48–2.43)  | 0.66 (0.63–0.69)  | 0.92 (0.90–0.94)                                   | 0.80 (0.77–0.83)   | 0.68 (0.62–0.75)  | 0.88 (0.85–0.91)  | 1.19 (0.45–3.18)   |
| Consolidated (>32.0% and TPC ≥100 for at least 4 years) | 2.05 (1.53–2.76)  | 0.53 (0.48–0.57)  | 0.84 (0.81–0.86)                                   | 0.61 (0.57–0.65)   | 0.53 (0.44–0.63)  | 0.88 (0.83–0.93)  | 0.62 (0.10–3.90)   |
| Number of observations                                  | 14166   | 15948   | 17118  | 17070  | 12528   | 17118   | 10776  |
| Number of municipalities                                | 2361  | 2658  | 2853   | 2845   | 2088  | 2853  | 1796   |

OR=odds ratio. RR=rate ratio. DTP=diphtheria, tetanus, pertussis. TPC=target population coverage. FHP=Family Health Programme. \*Estimated by logistic regression models adjusted for FHP coverage. †Estimated by negative binomial regression models adjusted for FHP coverage. ‡Estimated by negative binomial regression models adjusted for FHP coverage, income per person, proportion of municipality population eligible for BFP, proportion of individuals living in households with inadequate sanitation, proportion of individuals older than 15 years who are illiterate, and total fertility rate.

**Table 4: Fixed-effect models for associations between primary care indicators, rates of admission to hospital, and *Bolsa Familia* Programme (BFP) coverage**

in the models for overall under-5 mortality rate (RR 0.95; 95% CI 0.91–0.99).

Municipalities with adequate death and livebirth information showed a slightly lower socioeconomic status and a slightly higher BFP coverage than did those with inadequate information (data not shown). A sensitivity test, done by fitting the models with data from all Brazilian municipalities, showed slightly lower, but significant, effects of the two interventions: the effect on overall under-5 mortality of consolidated BFP coverage was RR 0.83 (95% CI 0.78–0.87) and of consolidated FHP coverage was 0.91 (0.87–0.94), while for under-5 diarrhoea mortality was 0.52 (0.41–0.66) for consolidated BFP and 0.65 (0.54–0.79) for consolidated

FHP. We identified much the same results for malnutrition and respiratory infection (data not shown).

## Discussion

The results of our study show that BFP had a significant role in reduction of under-5 mortality, overall and from poverty-related causes such as malnutrition and diarrhoea, in Brazilian municipalities from 2004–09. The effect was maintained even after the adjustment for socioeconomic covariables and FHP. The increase in BFP duration and in coverage of both the total and target populations strengthens the effect of the programme. The effect of the BFP was stronger when, with high municipality population coverage, full coverage of the

target population of poor families was maintained for 4 years or more. With regard to factors involved in the causal chain of mortality reduction, the BFP substantially reduced rates of under-5 admission to hospital and increased vaccination coverage and prenatal visits.

Because the BFP and FHP have been implemented on a large scale over the same period in the same areas in Brazil, we had a unique opportunity to explore their joint effects. The effectiveness of the FHP in reduction of overall infant and child mortality, reducing under-5 mortality resulting from diarrhoea and lower respiratory infections, increasing vaccination coverage, and reducing admissions to hospital has already been shown.<sup>11,12,22</sup> However, none of these studies included the effect of the BFP in their analyses.

Several studies worldwide, summarised in reviews,<sup>23–25</sup> showed that conditional cash transfer programmes had positive effects on nutritional status and health outcomes of enrolled children, through the increase in the use of preventive services, immunisation coverage, and promotion of healthy behaviours (panel<sup>23–25</sup>). Only an econometric study<sup>26</sup> assessed the effect of a conditional cash transfer programme on infant mortality; the investigators reported that the Mexican conditional cash transfer programme *Progres*a was able to reduce infant mortality in rural areas.<sup>26</sup> Our analysis—using a different statistical approach and different mortality outcomes, excluding municipalities

with inadequate vital information, using different coverage indicators, and studying the effect of BFP on the intermediate mechanisms of vaccination, prenatal care, and admissions to hospital—showed BFP to have an effect on childhood mortality.

The large magnitude of the effect of the BFP that we identified can be explained by the fact that the number of under-5 deaths from a small number of extremely poor families constitutes a high proportion of all under-5 deaths in the municipalities. The proportion reaches almost 100% for poverty-related causes of mortality, such as malnutrition or diarrhoea. A mathematical demonstration and a broader discussion of how this association operates are available in the appendix pp 5–6. Moreover, to understand how the relatively small amount of money provided by the BFP can have an effect on beneficiaries' health, it has to be remembered that this amount is proportional to the economic vulnerability of the families, and that the association between income and health is non-linear:<sup>28</sup> even a small amount of money, given to extremely poor families, can have an effect on child health, increasing survival.

BFP, like other conditional cash transfer programmes, can affect child survival through different mechanisms (figure), largely centred on income improvement and health conditions: an increased income can increase access to food and other health-related goods, and health-related conditions can improve access to health services.<sup>25</sup> A strong association exists between child undernutrition and child survival—as levels of child undernutrition increase so does the risk of death, especially from diarrhoea and measles.<sup>29</sup> Research has already shown that poor families enrolled in the BFP increased food expenditures and improved food security in their households.<sup>30</sup> Overall, Brazil has seen a substantial decrease in child undernutrition during the past decade, particularly among poor families.<sup>9</sup> The contribution of the BFP to this process has been shown in a few studies,<sup>13,14</sup> in which children from BFP beneficiary families were more likely to be better nourished than were those from non-beneficiary families. The money allowance from the BFP could likewise reduce the household poverty burden, improving living conditions and removing or reducing barriers to accessing health care, not only for children, but also the rest of the family.<sup>31</sup>

Another important explanation for the effect of the BFP on child survival is associated with the health-related conditions, which include prenatal care, postnatal care, and health and nutrition education activities for mothers, in addition to a regular vaccination schedule and routine check-ups for growth and development for children younger than 7 years. Maternal knowledge and education are some of the strongest determinants of child health, improving nutrition, hygiene practices, and care-seeking for illnesses.<sup>31</sup> Even when there is conflicting evidence as to whether monitoring of child growth is effective in itself, such monitoring can provide an entry point for preventive and curative health-care services and can reduce the scarcity of contact with the health system, an important

### Panel: Research in context

#### Systematic review

We searched PubMed, Science Direct, Popline, and Embase with the terms “conditional cash transfer” and “bolsa familia”. Searches were not restricted by language or date; the last search was done in June, 2012. We used the reference lists of selected reports to identify other relevant studies.

In a comprehensive systematic review<sup>23</sup> of the effect of conditional cash transfer programmes on general health outcomes, positive effects of such programmes were identified for child nutritional status and child morbidity. An increased use of general health care and preventive services in children and pregnant women was also reported.<sup>24</sup> A programme theory framework has been proposed to explain the effects of conditional cash transfer on health,<sup>25</sup> suggesting that the quality of the health services providing conditions is a crucial factor in the effectiveness of such programmes. With regard to the effect of conditional cash transfer programmes on mortality, we identified only one research report, which showed a reduction of infant mortality rates in Mexico, attributed to the effect of the conditional cash transfer programme *Progres*a,<sup>26</sup> probably because it increases access to health care for hard-to-reach segments of the population in both rural and urban areas.<sup>27</sup> However, limitations in the study design and the absence of analysis of some intermediate mechanisms that could explain this mortality reduction—including health-care supply—emphasised the necessity of a broader and more rigorous study of the effect on child mortality of a larger conditional cash transfer programme, such as *Bolsa Familia* in Brazil.

#### Interpretation

The results of our study show that a large-scale conditional cash transfer programme, combined with an effective primary health-care system, can strongly reduce childhood mortality, from poverty-related causes and overall. Mechanisms include effects on social determinants of health and increased use of preventive services in children and pregnant women through programme conditions.

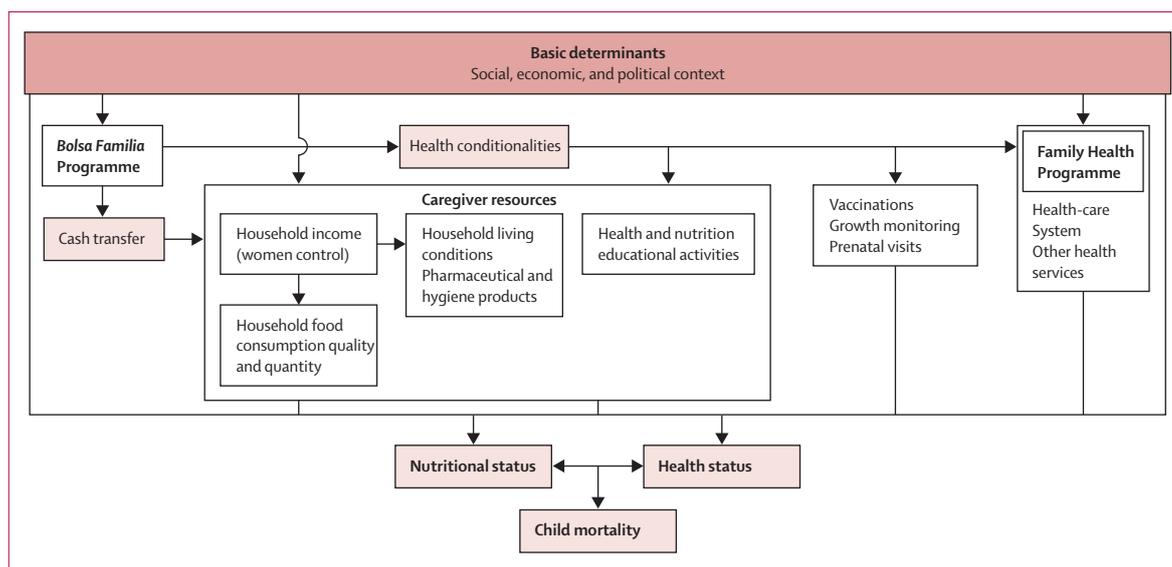


Figure: Mechanisms linking the Bolsa Familia Programme and the Family Health Programme to child nutritional and health outcomes

determinant of child survival in developing countries.<sup>32</sup> As shown in our study, the BFP increases prenatal care and vaccination coverage. These interventions are effective for prevention of child mortality.<sup>33</sup>

Even if implementation and coverage of the FHP is not affected by the presence of the BFP, according to the Ministry of Health, the FHP is the strategy of choice to help BFP beneficiaries meet health conditions, and to help them with their health needs.<sup>6</sup> Unlike the BFP, which has a specific target population, the FHP has the objective of covering all the population of the municipality, offering comprehensive primary health care free of charge.<sup>8</sup> When BFP beneficiaries are under an FHP catchment area, the FHP team has the formal responsibility to deliver all services linked to the conditions, and community health workers should undertake home visits and actively monitor the completion of conditions.<sup>6</sup> Compliance with health conditions depends on monitoring and on barriers to accessing services.<sup>1,25</sup> The FHP increases access to health care,<sup>15,34</sup> which, according to our results, strengthens the effect of the BFP on FHP beneficiaries compared with BFP beneficiaries who are assisted by traditional health facilities, which are generally more distant and do not undertake community involvement activities and home visits.<sup>3</sup>

We identified a strong effect of the BFP on rates of under-5 admissions to hospital, both overall and for specific causes, which could be explained by two different mechanisms: decreasing the incidence of the diseases by affecting social determinants of health; or by increasing early contacts with the health-care system, thereby reducing the number of severe cases of illness needing admission to hospital.<sup>25</sup> One of the strengths of our study is that we used a measurement of the intensity of the intervention (the coverage of the BFP) that is specifically linked to the population group that accounts for a large

proportion of the outcome (deaths from poverty-related causes), thus reducing the plausibility that the outcomes of interest are coming from the group of people not exposed to the intervention (ie, ecological fallacies).

Another strength of our study was the selection of municipalities with vital information of adequate quality, which improved the study's internal validity, although this selection might limit the generalisability of the results. However, the sensitivity analysis done with all Brazilian municipalities gave much the same effect estimates, suggesting that our findings are robust. In some of the models for selected causes of mortality, the number of observations varied for statistical reasons; municipalities with the same values for the outcome (in this case, 0 deaths) over the entire 6-year period were not included in the model fitting because of a limitation of the fixed-effects model algorithms.<sup>19,20</sup> However, by comparing the covariate values of the municipalities included in each model with those that were excluded, we identified much the same values, and the estimates of the random effects models (which included all 2853 municipalities in the model fitting) did not affect the sign, significance, or main conclusions reached with the fixed-effects models; the random-effects RR for consolidated BFP coverage was 0.75 for mortality from diarrhoea and 0.61 for mortality from malnutrition. Possible bias introduced by use of crude interpolation rather than more complex estimation techniques was limited by variable categorisation, which can reduce fluctuations that are artificially introduced by the method.

We did not include a variable representing time in the models because the mortality RR, comparing two or more groups of coverage exposed to the same mortality time trend, allowed us to control for secular trends.<sup>11,12</sup> The introduction of a time variable in the models would have constituted an over-specification problem, as confirmed

by the sensitivity analyses. The fact that these models were not affected by secular mortality trends was suggested by the estimates of the effect of BFP and FHP on under-5 mortality attributed to external causes: although mortality from external causes was decreasing during the study period, neither programme showed an effect of reduction on mortality from such causes. One limitation of the study was that fixed-effects models can control only for selection bias associated with unmeasured time-constant characteristics of the municipalities.<sup>21</sup> However, we used a wide set of covariates and showed no effects of either programme on mortality from external causes, suggesting that other possible sources of selection bias or confounding were controlled.

The results of our study provide evidence that a multisectoral approach, combining a large-scale conditional cash transfer programme, with the potential to act on important social health determinants, and effective primary health care, capable of attending basic health demands of the same population and of attending conditions imposed by the conditional cash transfer programme, can substantially reduce childhood mortality from poverty-related causes in a large middle-income country such as Brazil.

#### Contributors

MLB, DR, and RA designed the original study in discussion with RP-S. DR collected data. DR and CATS did the data analysis and DR wrote a first draft of the report. All authors contributed to data interpretation and to the review of the report.

#### Conflicts of interest

We declare that we have no conflicts of interest.

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#### References

- World Bank. Conditional cash transfer. A World Bank policy research report. Washington: World Bank, 2009.
- Handa S, Davis B. The experience of conditional cash transfers in Latin America and the Caribbean. *Development Policy Rev* 2006; **24**: 513–536.
- Lindert K, Linder A, Hobbs J, Briere B. The nuts and bolts of Brazil's *Bolsa Família* Program: implementing conditional cash transfers in a decentralized context. <http://siteresources.worldbank.org/INTLACREGTOPLABSOCPRO/Resources/BRBolsaFamiliaDiscussionPaper.pdf> (accessed April 8, 2013).
- Ministerio do Desenvolvimento Social e Combate a Fome. Matriz de Informação Social. [http://aplicacoes.mds.gov.br/sagi/mi2007/tabelas/mi\\_social.php](http://aplicacoes.mds.gov.br/sagi/mi2007/tabelas/mi_social.php) (accessed April 8, 2013).
- Ministerio do Desenvolvimento Social. Bolsa Família. <http://www.mds.gov.br/bolsafamilia/> (accessed April 8, 2013).
- Ministerio da Saúde. Manual de orientações sobre o Bolsa Família na Saúde. 2009. [http://bvsmms.saude.gov.br/bvs/publicacoes/manual\\_orientacao\\_sobre\\_bolsa\\_familia.pdf](http://bvsmms.saude.gov.br/bvs/publicacoes/manual_orientacao_sobre_bolsa_familia.pdf) (accessed April 8, 2013).
- Ministerio da Saúde. Departamento de Informática do SUS. DATASUS. <http://www2.datasus.gov.br/datasus/index.php> (accessed April 8, 2013).
- Ministerio da Saúde. Guia pratico do programa da Saude da Familia. Brasilia: Ministerio da Saude, 2001.
- Victora CG, Aquino EM, do Carmo Leal M, Monteiro CA, Barros FC, Szwarcwald CL. Maternal and child health in Brazil: progress and challenges. *Lancet* 2011; **377**: 1863–76.
- UNICEF. Brazil, statistics. [http://www.unicef.org/infobycountry/brazil\\_statistics.html](http://www.unicef.org/infobycountry/brazil_statistics.html) (accessed April 8, 2013).
- Aquino R, Oliveira NF, Barreto ML. Impact of the Family Health Program on infant mortality in Brazilian municipalities. *Am J Public Health* 2009; **99**: 87–93.
- Rasella D, Aquino R, Barreto ML. Reducing childhood mortality from diarrhea and lower respiratory tract infections in Brazil. *Pediatrics* 2010; **126**: e534–40.
- Paes-Sousa R, Santos LM, Miazaki ES. Effects of a conditional cash transfer programme on child nutrition in Brazil. *Bull World Health Organ* 2011; **89**: 496–503.
- Paes-Sousa R, Quiroga J. Programas de transferência de renda e redução da pobreza e das desigualdades sociais no Brasil. Brazil: Ministerio da Saude, 2011.
- Rasella D, Aquino R, Barreto ML. Impact of the Family Health Program on the quality of vital information and reduction of child unattended deaths in Brazil: an ecological longitudinal study. *BMC Public Health* 2010; **10**: 380.
- Andrade CLT, Szwarcwald CL. Sociospatial inequalities in the adequacy of Ministry of Health data on births and deaths at the municipal level in Brazil, 2000–02. *Cad Saude Publica* 2007; **23**: 1207–16 (in Portuguese).
- WHO. The Global Burden of Disease: 2004 update. [www.who.int/healthinfo/global\\_burden\\_disease/2004\\_report\\_update/en/index.html](http://www.who.int/healthinfo/global_burden_disease/2004_report_update/en/index.html) (accessed April 8, 2013).
- IBGE. Instituto Brasileiro de Geografia e Estatística. <http://www.ibge.gov.br> (accessed April 8, 2013).
- Hilbe JM. Negative binomial regression. Cambridge, UK: Cambridge University Press, 2007.
- Wooldridge JM. Introductory Econometrics, a modern approach, 3rd edn. Cincinatti: South-Western College Publishers, 2005.
- Shahidur RK, Koolwal GB, Samad HA. Handbook on impact evaluation: quantitative methods and practices. Washington DC: World Bank Publications, 2010.
- Macinko J, de Oliveira VB, Turci MA, Guanais FC, Bonolo PF, Lima-Costa MF. The influence of primary care and hospital supply on ambulatory care-sensitive hospitalizations among adults in Brazil, 1999–2007. *Am J Public Health* 2011; **101**: 1963–70.
- Lagarde M, Haines A, Palmer N. The impact of conditional cash transfers on health outcomes and use of health services in low and middle income countries. *Cochrane Database Syst Rev* 2009; **4**: CD008137.
- Ranganathan M, Lagarde M. Promoting healthy behaviours and improving health outcomes in low and middle income countries: a review of the impact of conditional cash transfer programmes. *Prev Med* 2012; **55** (suppl): S95–105.
- Gaarder MM, Glassman A, Todd JE. Conditional cash transfers and health: unpacking the causal chain. *J Develop Effect* 2010; **2**: 6–50.
- Barham T. A healthier start: The effect of conditional cash transfers on neonatal and infant mortality in rural Mexico. *J Development Econ* 2011; **94**: 74–85.
- Knaut FM, González-Pier E, Gómez-Dantés O, et al. The quest for universal health coverage: achieving social protection for all in Mexico. *Lancet* 2012; **380**: 1259–79.
- Ecob R, Smith GD. Income and health: what is the nature of the relationship? *Soc Sci Med* 1999; **48**: 693–705.
- Black RE, Allen LH, Bhutta ZA, et al, for the Maternal and Child Undernutrition Study Group. Maternal and child undernutrition: global and regional exposures and health consequences. *Lancet* 2008; **371**: 243–60.
- De Bem Lignani J, Sichieri R, Burlandy L, Salles-Costa R. Changes in food consumption among the *Programa Bolsa Família* participant families in Brazil. *Public Health Nutr* 2011; **14**: 785–92.
- Wagstaff A, Bustreo F, Bryce J, Claeson M, for the WHO-World Bank Child Health and Poverty Working Group. Child health: reaching the poor. *Am J Public Health* 2004; **94**: 726–36.
- Gwatkin D, Habicht JP, Bryce J, et al, for the Multi-Country Evaluation of IMCI Study Group. Reducing child mortality: can public health deliver? *Lancet* 2003; **362**: 159–64.
- Jones G, Steketee RW, Black RE, Bhutta ZA, Morris SS, for the Bellagio Child Survival Study Group. How many child deaths can we prevent this year? *Lancet* 2003; **362**: 65–71.
- Macinko J, Lima Costa MF. Access to, use of and satisfaction with health services among adults enrolled in Brazil's Family Health Strategy: evidence from the 2008 National Household Survey. *Trop Med Int Health* 2012; **17**: 36–42.